Partial Autocorrelation Function (PACF).

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Econometrics

We consider the time series AR(p) in the form:

$$X_t = \phi_1 X_{t-1} + \phi_2 X_{t-2} + \ldots + \phi_p X_{t-p} + \epsilon_t,$$

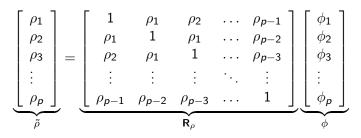
where $\phi_1, \phi_2, \ldots, \phi_p$ are unknown parameters. Simple example shows that AR(p) may be stationary or not:

- the model AR(1) in the form $X_t = X_{t-1} + \epsilon_t$ is not stationary;
- the model AR(1) in the form X_t = φX_{t−1} + ε_t is stationary if and only if φ ∈ (−1, 1);

- Assume that our model AR(p) is stationary.
- We find the preliminary estimator of φ₁, φ₂,..., φ_p using Yule-Walker formula.

Yule-Walker formula

We have derived the Yule-Walker formula:



Remarks:

- The matrix R_ρ of correlations of the vector (X₁, X₂,..., X_n) is positively definite, hence nonsingular;
- **2** Hence there exists an inverse matrix \mathbf{R}_{ρ}^{-1} ;
- So The vector ϕ is uniquely determined $\phi = \mathbf{R}_{\rho}^{-1}\tilde{\rho}$, i.e. only such a vector determines a stationary model type AR(p);

The preliminary estimation of parameters of stationary time series AR(p)

Since we observe $X_1, X_2, ..., X_n$ we can obtain the sampling autocorrelations:

$$\hat{\rho}_h = \frac{\hat{\gamma}_h}{\hat{\gamma}_0},$$

where for $h \ge 0$ we have

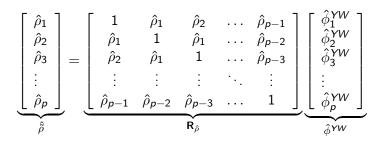
$$\hat{\gamma}_h = \frac{1}{n} \sum_{t=1}^{n-h} (X_t - \bar{X}) (X_{t+h} - \bar{X})$$

and

$$\bar{X} = \frac{1}{n} \sum_{t=1}^{n} X_t.$$

The preliminary estimation of parameters of stationary time series AR(p)

Substituting ρ_h by its estimator $\hat{\rho}_h$ the Yule-Walker furmula provide



Multiplying both sides by the corresponding inverse matrix we have the **preliminary estimator** of $\phi_1, \phi_2, \ldots, \phi_p$,

$$\hat{\phi}^{YW} = \mathbf{R}_{\hat{\rho}}^{-1} * \hat{\tilde{\rho}}.$$

The preliminary estimation of parameters of stationary time series AR(p)

The $\hat{\phi}^{YW}$ is a preliminary estimator of ϕ and cannot be a final estimation due to the following reasons:

- The model AR(p) is not always stationary, in practice the assumption of stationarity is not reasonal;
- Even if the model AR(p) is stationary the value of φ^{YW} is sensitive to rounding errors, especially if the parameters of AR(p) are close to the *edge* of stationarity;
- But because we usually are not aware the value of p, φ̂^{YW} is usefull as a preliminary estimator, which we apply for determining p (the details later);
- The final estimator of ϕ will be determined using the maximum likelihood method method.

Yule Walker formula. What else?

If X_t type AR(p) in the form

$$X_t = \phi_1 X_{t-1} + \phi_2 X_{t-2} + \ldots + \phi_p X_{t-p} + \epsilon_t$$

can be embedded into AR(p+1) as well:

$$X_{t} = \phi_{1}X_{t-1} + \phi_{2}X_{t-2} + \ldots + \phi_{p}X_{t-p} + 0 * X_{t-(p+1)} + \epsilon_{t}$$

and consequently in the arbitrary AR(p + m)

$$X_{t} = \phi_{1}X_{t-1} + \phi_{2}X_{t-2} + \ldots + \phi_{p}X_{t-p} + \phi_{p+1}X_{t-(p+1)} + \phi_{p+2}X_{t-(p+2)} + \ldots + \phi_{p+m}X_{t-(p+m)} + \phi_{p+1} = \phi_{p+1} = \phi_{p+1} = \phi_{p+m} = 0 + \epsilon_{t}.$$

Now we can say X_t is AR(p) if

$$\phi_{p} \neq 0 \quad \text{and} \quad \phi_{p+1} = \phi_{p+2} = \ldots = 0,$$

hence ϕ_p is the last non-zero coefficient;

- If X_t is really type AR(p), but we compute Yule-Walker formula AR(p+m) for some m > 0 one can obtain:
 - $\hat{\phi}_{p}^{YW}$ relatively far from 0, and all next coefficients $\hat{\phi}_{j}^{YW}$ (j > p) should be relatively close to 0:
 - intuitively, whe should obtain

$$|\hat{\phi}_p^{YW}| > \text{a threshold }, \hat{\phi}_{p+1}^{YW} \approx 0, \hat{\phi}_{p+2}^{YW} \approx 0, \dots, \text{ and } \hat{\phi}_{p+m}^{YW} \approx 0,$$

and the values $\hat{\phi}_1^{YW}, \hat{\phi}_2^{YW}, \dots, \hat{\phi}_{p-1}^{YW}$ are less informative for our analysis.

For using Yule-Walker formula:

- We presume X_t is stationary and is in type of AR(p), but we are not aware the size of p;
- 2 We are able to predict the real value of p;
- Having the prediction p, we pass to further analysis with AR(p), and we further ignore the value $\hat{\phi}^{YW}$ (it is a preliminary estimator only);
- We find estimators of \(\phi_1, \phi_2, \ldots, \phi_p\) using maximum likelihood method (it will be later).

We construct a Partial Autocorrelation Function (PACF)

- We base on time series X_1, X_2, \ldots, X_n ;
- We apply algorithm of Yule-Walker n 1 times successively extending the model from AR(1) through AR(2),AR(3),...,AR(p).
- To distinguish Yule-Walker estimators obtained by distinct models AR(p) we denote

$$\hat{\phi}_{1,p}^{YW} \approx \phi_1, \hat{\phi}_{2,p}^{YW} \approx \phi_2, \dots$$
 , and $\hat{\phi}_{p,p}^{YW} \approx \phi_p$.

Partial Autocorrelation Function (PACF) - algorithm

1 If p = 1, the Yule-Walker formula for AR(1) provides

$$\hat{\phi}_{1,1}^{YW} = \hat{\rho}_1 \approx \phi_1;$$

3 If p = 2, the Yule-Walker formula for AR(2) provides

$$\hat{\phi}_{1,2}^{YW} \approx \phi_1$$
 and $\hat{\phi}_{2,2}^{YW} \approx \phi_2$;

③ continue this procedure for any p satisfying $1 \le p \le n-1$:

$$\hat{\phi}_{1,p}^{YW} \approx \phi_1, \hat{\phi}_{2,p}^{YW} \approx \phi_2, \quad , \dots, \text{and} \quad \hat{\phi}_{p,p}^{YW} \approx \phi_p.$$

• The function $p \mapsto \hat{\phi}_{p,p}^{YW}$ is called the **Partial Autocorrelation** Function.

PACF - application

Assuming that the time series X_t is type AR(p) we compute $\hat{\phi}_{p,p}^{YW}$:

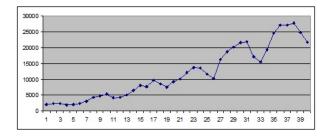
- in next units of lectures we put significance thresholds for $\hat{\phi}_{p,p}^{YW}$;
- if the moduli of \$\hfip\delta_{p,p}^{YW}\$ extends this threshold, we say \$X_{t-p}\$ is insignificant in the computing \$X_t\$;
- to accept that the truth model is p one must have
 - the value $\hat{\phi}_{p,p}^{YW}$ is significant;
 - for of k > p most of the values of $\hat{\phi}_{k,k}^{YW}$ are insignificant, and we allow sparse elements of $\hat{\phi}_{k,k}^{YW}$ to be significant provided that they are close to significance threshold.

COVID in Poland 1.10.2020-9.11.2020

Unfortunately there is no country free of corona-virus COVID-19. The daily number of reported infections in Poland from (1.10.2020) until yesterday (9.11.2020) is reported in the table in the next page. Let X_t be the time series which means the number of infected people in day t, where t the natural index of days in considered period.

The number of reported infections in Poland during 1.10-9.11.2020.

| Day | 1-4.10 | 5-11.10 | 12-18.10 | 19-25.10 | 26.10-1.11 | 2-8.11 | 9.11 |
|-----------|--------|---------|----------|----------|------------|--------|-------|
| Monday | | 2006 | 4324 | 7482 | 10241 | 15578 | 21713 |
| Tuesday | | 2236 | 5068 | 9291 | 16300 | 19364 | |
| Wednesday | | 3003 | 6526 | 10040 | 18820 | 24692 | |
| Thursday | 1967 | 4280 | 8099 | 12107 | 20156 | 27143 | |
| Friday | 2292 | 4739 | 7705 | 13632 | 21629 | 27086 | |
| Saturday | 2367 | 5300 | 9622 | 13628 | 21897 | 27876 | |
| Sunday | 1934 | 4178 | 8536 | 11742 | 17171 | 24785 | |



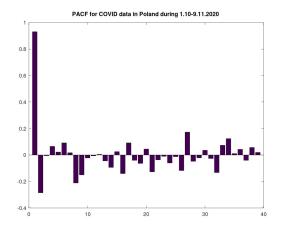
The plot of the trajectory of X_t in the following picture:

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Analysis in Octave yields us the following bar plot of PACF for the time series X_t . The height of the bar at p means the value $\hat{\phi}_{p,p}^{YW}$.



The dominating value $\hat{\phi}_{1,1}^{YW}$ over all leads us to suppose that the model is AR(1).

Preliminary corollaries

- Based on analysis of PACF only we may presume X_t has AR(1) form;
- Yule-Walker formula yields $\phi_{1,1}^{YW} = 0.9309$ hence the model is

$$X_t = 0.9309 * X_{t-1} + \epsilon_t;$$

- But Yule-Walker formula always assume stationarity of X_t which is not always acceptable;
- This model is not final and will be corrected using the maximal likelihood method, where we relax the assumption of stationarity and include a constant value:

$$X_t = \mu + \phi X_{t-1} + \epsilon_t;$$

• With the calculus we have

$$X_t = 11866.6 + 0.972640 * X_{t-1} + \epsilon_t;$$

Remark

Both models should be diagnostically verified and at this stage this model is far to be believable. For example

- we have not compared PACF with the thresholds of significance;
- we have not verified the normality of residua, equality of variance, missing of correlation;
- we have not taken into consideration other phenomena like seasonality (e.g. on Monday we have always less infected people than on Wednesday);
- in fact we do not know whether dependence of infected today and tomorrow have linear form.

Because of that, please do not believe this model. It is very preliminary.

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Preliminary corollaries

But if the following model

 $X_t = 11866.6 + 0.972640 * X_{t-1} + \epsilon_t;$

is correct, then the expectation is

 $EX_t = 11866.6 + 0.972640 * EX_{t-1}$.

Then,

- yesterday we had 21713 infected people;
- tomorrow we expect

 $EX_{41} = 11866.6 + 0.972640 * 21713 = 32985.53;$

• and the expectation will increase until it reaches per day

$$\lim_{t \to \infty} E_t = \frac{11866.6}{1 - 0.972640} \approx 433720$$

• According to this model, there is no evidence that the epidemic is going to slow down, quite the contrary.